RESEARCH ARTICLE

Quantitative Assessment the Relationship between p21 rs1059234 Polymorphism and Cancer Risk

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Abstract

p21 is a cyclin-dependent kinase inhibitor, which can arrest cell proliferation and serve as a tumor suppressor. Though many studies were published to assess the relationship between p21 rs1059234 polymorphism and various cancer risks, there was no definite conclusion on this association. To derive a more precise quantitative assessment of the relationship, a large scale meta-analysis of 5,963 cases and 8,405 controls from 16 eligible published case–control studies was performed. Our analysis suggested that rs1059234 was not associated with the integral cancer risk for both dominant model [(T/T+C/T) *vs* C/C, OR=1.00,95% CI: 0.84-1.18] and recessive model [T/T *vs* (C/C+C/T), OR=1.03,95% CI: 0.93-1.15)]. However, further stratified analysis showed rs1059234 was greatly associated with the risk of squamous cell carcinoma of head and neck (SCCHN). Thus, larger scale primary studies are still required to further evaluate the interaction of p21 rs1059234 polymorphism and cancer risk in specific cancer subtypes.

Keywords: p21 - cancer risk - polymorphism - meta-analysis

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Introduction

It has been suggested that environmental and genetic factors may affect the individual's susceptibility to cancer(Derynck et al., 2001). An important gene identified as cancer susceptibility one is p21 (also known as CDKN1A), a member of the Cip/Kip family of cyclindependent kinase (CDK) inhibitors. Expression of p21 is up-regulated by wildtype p53 in response to DNA damage to induce cell cycle arrest at the G1 checkpoint (Xiong et al., 1993; Sherr, 1996). p21 can exerts tumorsuppressive effects by inhibiting PCNA-dependent DNA replication and mismatch repair(Li et al., 1994; Waga et al., 1997). Somatic mutations in the p21 gene are rare in human malignancies(Roninson, 2002). However, reduced p21 expression in tumors has been associated with poor prognosis in humans (Jiang et al., 1997; Wakasugi et al., 1997). Therefore, genetic polymorphisms in p21 may modulate its expression and thereby affect carcinogenesis.

p21 polymorphism rs1059234 (C70T) locates within the 3' untranslated region of p21 gene, causes a single C-to-T substitution 20 nt downstream of the stop codon at exon 3 (http://egp.gs.washington.edu). This polymorphic variant identified was thought to alter p21 function and maybe functionally associated with cancer susceptibility.

Analysis of case-control studies is the most prevalent method of investigating the association between a disease and a specific gene polymorphism. Thus far, a number of studies have reported the role of p21 rs1059234 polymorphism in cancer risk (Li et al., 2005; Lei et al., 2010), but the results remain conflicting (Sivonova et al., 2013; Yin et al., 2015), partially because of the relatively small sample size in each of the published studies. Therefore, here we performed a large scale meta-analysis of all the published studies to derive a more precise quantitative assessment of the association between p21 rs1059234 polymorphism and the cancer risk.

Materials and Methods

Selection of studies

All of the case-control studies were identified by a computerized literature search of the PubMed, Web of Science, EBSCO, and CGEMS database (prior to March 2015) using the following words and terms: "p21", "CDKN1A", "polymorphism", and "cancer". References of the retrieved publications were also screened. Studies had to be based on an unrelated case-control design, so pedigree data were excluded. The following basic data were collected from the studies: first authors, journals, year of publications, cancer subtypes and ethnicity of the population.

Statistical analysis

For each study, the OR was first calculated to assess the association between the polymorphisms and the disease

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Yongsheng Huang et al

in table 1. In meta-analysis, we examined the association between p21 rs1059234 polymorphism and the risk of cancer using recessive [T/T vs (C/C + C/T)] and dominant [(T/T + C/T) vs C/C] genetic models. There are three widely used methods of meta-analysis for dichotomous outcomes: two fixed effects methods (Mantel-Haenszel's method and Peto's method), which assume that studies are sampled from populations with the same effect size, making an adjustment to the study weights according to the in-study variance; and one random effects method (DerSimonian and Laird's method), which assumes that studies are taken from populations with varying effect sizes, calculating the study weights both from the in-study and between-study variance, considering the extent of variation, or heterogeneity. In our study, both Mantel-Haenszel's fixed effects method and DerSimonian and Laird's random effects method were used in Stata 10.0 software. A chi-square-based Q-statistic test was performed to evaluate the between-study heterogeneity of the studies. If P < 0.10, the between-study heterogeneity was considered to be significant, we chose the randomeffects model to calculate the OR. Otherwise, when $P \ge 0.10$, the between-study heterogeneity was not significant, then the fixed-effects model was suitable. In the absence of between-study heterogeneity, the two methods yield similar results. In order to make a clear comparison, we present the OR of both the random-effects model and fixed-effects model for every meta-analysis. A pooled OR obtained by meta-analysis was used to give a more reasonable evaluation of the association. A Z test was performed to determine the significance of the pooled OR ($P \leq 0.05$ suggests a significant OR). Funnel plots were used to access publication bias by the method of Egger's regression test. A T test was performed to determine the significance of the asymmetry. An asymmetric plot suggested possible publication bias ($P \ge 0.05$ suggests no bias). Hardy-Weinberg equilibrium was tested by the

Chisquare test based on a program (http://www.ihg.gsf. de/cgi-bin/hw/hwa1.pl). Analyses were performed by Stata10.0 software.

Results

Study characteristics

There are 16 studies (5,963 cases and 8,405 controls) analyzing the relation of p21 rs1059234 polymorphism and the risk of cancer. Each subpopulation in these articles was treated as a separate study in our meta-analysis. All the studies were published from Year 2005 to 2015. Populations were divided into different cancer subtypes and ethnic categories. Table 1 shows the details of the cases and controls in the included studies, together with the ORs we calculated to make a primary evaluation. Table 2 is the summary of the meta-analysis of case-control studies examining the association between p21 rs1059234 polymorphism and cancer risk, with the comparison between different cancer subtypes and different ethnicities.

Main results

For each study we investigated the association between p21 rs1059234 polymorphism and cancer risk, assuming different inheritance models of the C70T allele. Overall, when all the eligible studies were pooled into the meta-analysis, no associations between p21 rs1059234 polymorphism and cancer susceptibility were observed in all genetic models. No significant associations were found for T/T vs C/C (OR=1.02; 95% CI: 0.83-1.26; P=0.010 for heterogeneity), C/T vs C/C (OR=0.98; 95% CI: 0.82-1.16; P=0.000 for heterogeneity), T/T+C/T vs C/C (OR=1.00; 95% CI: 0.84-1.18; P=0.000 for heterogeneity) and T/T vs C/T+T/T (OR=1.03; 95% CI: 0.93-1.15; P=0.135 for heterogeneity)(Table 2). However, subgroup analyses by cancer type showed rs1059234 polymorphism might associate with the risk of SCCHN for T/T+C/T vs C/C

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Genotype distribution of p21 rs1059234 polymorphism											
				Case Control		OR (95%CI)		HWE			
Author	Cancer subtype	Population	C/C	C/T	T/T	C/C	C/T	T/T	T/T vs	(T/T + C/T)	(control
		-							(C/C + C/T)	vs C/C	P value)
Li et al., 2005	SCCHN	Caucasian	596	110	6	1080	136	6	1.72(0.55-5.36)	1.48(1.14-1.93)	0.445
Wu et al., 2006	bladder cancer	Caucasian	513	86	3	506	82	3	0.98(0.20-4.88)	1.03(0.75-1.42)	0.8692
Ma et al., 2006	Breast cancer	Asian	87	211	70	129	253	85	1.06(0.74-1.50)	1.23(0.90-1.69)	0.044
Guo et al., 2006	ESCC	Asian	94	154	51	166	221	50	1.59(1.04-2.43)	1.34(0.98-1.82)	0.0655
Guo et al., 2006	gastric cancer	Asian	95	121	50	166	221	50	1.79(1.17-2.74)	1.10(0.80-1.51)	0.0655
Driver et al., 2008	prostate cancer	Caucasian	167	100).Q	181	39	1	1.19(0.07-19.14)	0.51(0.29-0.92)	0.4711
Polakova et al., 2009	Colorectal Cancer	Caucasian	534	69	4	520	6.3	1	4.04(0)45-36.25)	0.79(0.57-1.10)	0.1603
Lei et al., 2010)	SCCHN	Caucasian	93	25	2	1009	139	14	1.39(0 31-6. 29) ³	1.91(1.21-3.04)	0.0004
Liu et al., 2010	Colorectal Cancer	Asian	100	197	76	223	438	177	0.96(0.71-1.29)	0.99(0.75-1.30)	0.1603
Taghavi et al., 2010	ESCC	Caucasian	99	275	o.Ø	82	18	0	Excluded	0.61(8 54 9 -0.83)	0.3227
Wang et al., 2012	cervical cancer	Asian	131	160	102	102	221	111	1.02(0.75-1.39)	1.05(0.80-1.38)	0.6942
Liu et al., 2013	hepatocellular cancer	Asian	134	224	118	153	56.3	118	14.648 (0.85-1.52)	2.15(1.23-3.76)	0.5493
Carvalho et al., 2013	retinoblastoma	Mixed (Braz)	90	49	2	95	23	2	0.85(0.12-6.12)	0.88(0.60-1.13)	0.6599
Sivonova et al., 2013	prostate cancer	Caucasian	104	14u	J.Y	108	22	0	Excluded 54.2	0.78(31.3-1.05)	0.2919
Zheng et al., 2014	ESCC	Asian	172	321	107	170	340	141	<u>0.79(0.</u> 59-1.04)	0.56(0.39-0.81)	0.2342
Shao et al., 2014	gastric cancer	Asian	99	158	56	154	301	126	0.79(0.55-1.12)	1.24(0.64-2.41)	0.3527
Yin et al., 2015	Endometrial Cancer	Asian	88	110	565	69	165	81	0.95(0.65-1.38)	0.66(0.32-1.36)	0.3831
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 Table 1. Characteristics of Studies Included in the Meta-Analysis

*OR: Odds ratio, CI: confidence interval, HWE: Hardy–Weinberg equilibrium, SCC HY: Squam **38.0** carcinoma of the head and neck, ESCC: esophageal squamous cell carcinoma

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4436 Asian Pacific Journal of Cancer Prevention, Vol 16, 2015



33.1

12.8

51.1

30.0

30.0

30.0

None

OR(95%CI)	T/T vs C/C	C/T vs C/C	T/T <i>vs</i> (C/C+C/T)	(T/T+C/T) <i>vs</i> T/T
Studies				
All of studies	1.02(0.83-1.26)	0.98(0.82-1.16)	1.03(0.93-1.15)	1.00(0.84-1.18)
All of SCCHN studies	1.72(0.70-4.23)	1.56(1.23-1.98)	1.60(0.65-3.92)	1.57(1.25-1.98)
All of ESCCstudies	1.14(0.48-2.70)	1.06(0.87-1.28)	1.10(0.55-2.19)	1.05(0.87-1.27)
All of colorectal Cancer studies	1.00(0.71-1.43)	0.89(0.71-1.11)	0.99(0.74-1.33)	0.90(0.73-1.11)
All of gastric cancerstudies	1.09(0.44-2.71)	0.88(0.70-1.11)	1.17(0.53-2.63)	0.92(0.74-1.14)
All of Caucasian studies	1.67(0.82-3.37)	1.02(0.74-1.41)	1.61(0.80-3.25)	1.03(0.75-1.42)
All of Asian studies	0.98(0.78–1.25)	0.89(0.75-1.07)	1.05(0.89–1.24)	0.92(0.77-1.10)
A) _{Study -}	Odds ratio (95% Cl)	B) _{Study -}	Í	Odds ratio (95% CI) % Weight

 Table 2. Summary of the Meta-analysis of Case-control Studies Examining the Association between p21 rs1059234

 Polymorphism and Cancer Risk



Figure 1. Forest Plot of Cancer Risk Associated with the p21 rs1059234 Polymorphism. The squares and horizontal lines correspond to the study-specific OR and 95%CI. The area of the squares reflects the weight (inverse of the variance). The diamond represents the summary OR and 95%CI. A) recessive genetic model [T/T vs (T/T+C/T)] B) dominant genetic models [(T/T+C/T) vs C/C]





(OR=1.57; 95% CI: 1.25-1.98; *P*=0.342 for heterogeneity) and T/T *vs* C/T+C/C (OR=1.60; 95% CI: 0.65-3.92; *P*=0.823 for heterogeneity) (Table 2).

Sensitivity analyses and publication bias

The results suggested the influences of the individual data set to the pooled ORs are all not significant. Funnel plots and Egger's test were performed to assess publication bias (Figure 2). The data suggested that there is no publication bias for the comparison of rs1059234 polymorphism C70T allele T vs allele C (t=1.03, P=0.318).

Discussion

Cell cycle control is crucial for normal cell growth

and differentiation and is regulated by cyclin-dependent kinases (CDKs). p21 is one of the universal inhibitors of cyclin-dependent kinases (CDK2, CDK3, CDK4, and CDK6)(Gartel and Tyner, 2002). It was initially discovered as a p53-target gene, but also has been suggested to play a role as a tumor suppressor in other cellular pathways including TGF- β and Wnt (Englert et al., 1997; Suzuki et al., 2012). Given the functional importance of p21 in carcinogenesis, genetic alteration of p21 could be associated with cancer risk.

So far, the functional role of the p21 rs1059234 variant has not yet to be well interpreted, several published clinic studies reported this variant was at increased risk of developing various cancer(Li et al., 2005; Lei et al., 2010; Liu et al., 2013). However, a number of published clinic studies reported this variant was not involved in the risk of cancer (Wu et al., 2006; Zheng et al., 2014; Yin et al., 2015). These conflicting studies based their conclusions on a small number of samples, so a metaanalysis of all available studies will help to establish a more convincing result. From our meta-analysis, p21 rs1059234 polymorphism in the combined population did not associate with cancer risk. There is no publication bias among the total studies. However, in the stratified analysis by ethnicity and subtype of cancer, significant association between p21 rs1059234 polymorphism and the risk of SCCHN was detected.

In conclusion, the research of the relationship of p21 rs1059234 polymorphism and cancer is very popular but remain conflicting at present. Our meta-analysis suggested that under recessive, dominant and other genetic models,

Yongsheng Huang et al

the p21 rs1059234 polymorphism did not associated with integral cancer risk. However, the studies included in the subgroups analysis are still limited and the results are sensitive to study selection. Since p21 also has a dual role can assume both pro- or anti-apoptotic functions in response to anti-tumor agents, depending on the cell type and context (Liu et al., 2003; Gartel, 2005). More comparative studies are needed to evaluate interactions of p21 rs1059234 polymorphism and cancer risk in specific cancer subtypes, especially in SCCHN.

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